AD Model Builder introduction course

## **Estimating uncertainties with AD Model Builder**

AD Model Builder foundation

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### Important to supply uncertainty estimates?

- I can think of three reasons people give for not to presenting uncertainty estimates:
  - 1. Estimates are so precise that we don't need to worry about uncertainties
  - 2. Estimates are so uncertain that we prefer not to show uncertainties
  - 3. We don't know how to estimate the uncertainties, only the estimates themselves.
- The uncertainty estimates are sort of needed in the first case too, and in the last two cases the estimates themselves should be skeptically evaluated.



## **AD Model Builder offers three approaches**

- Hessian based
- Profile likelihood
- MCMC

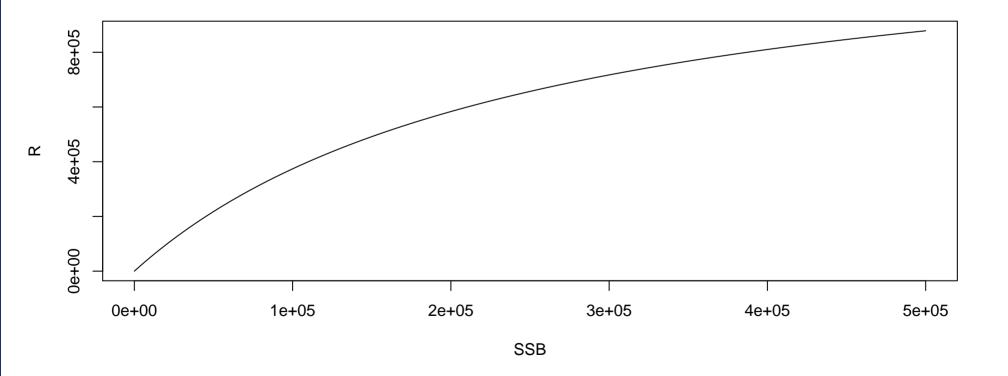


#### The Beverton-Holt stock recruitment example

• The Beverton-Holt model can be written (slightly re-parametrized) as:

 $\log R = \log(a) + \log(\mathsf{ssb}) - \log(1 + \exp(\log(b))\mathsf{ssb})$ 

• From a number of observed pairs {ssb<sub>i</sub>, log R<sub>i</sub>} i = 1...n we want to estimate the model parameters log(a) and log(b)



### Maximum likelihood estimator and Hessian

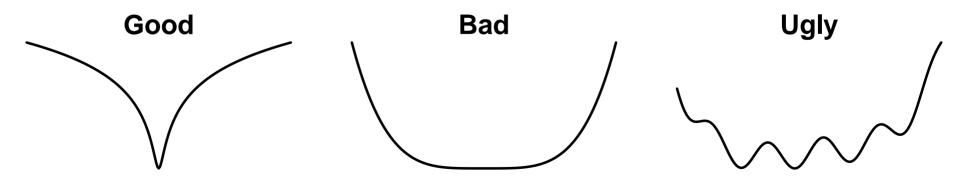
• A sensible estimate of the model parameters is to choose the values that maximize the likelihood for the actual observations.

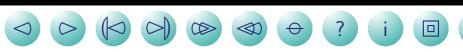
$$\widehat{\theta} = \mathop{\rm argmin}_{\theta} \ell(y|\theta)$$

• The curvature of the negative log likelihood function gives an estimate of the maximum likelihood estimator:

$$\widehat{\operatorname{var}(\widehat{\theta})} = \left( \frac{\partial^2 \ell(y|\theta)}{\partial \theta^2} \Big|_{\theta = \widehat{\theta}} \right)^{-1}$$

- The matrix  $\mathcal{H}(\widehat{\theta}) = \left(\frac{\partial^2 \ell(y|\theta)}{\partial \theta^2}\Big|_{\theta = \widehat{\theta}}\right)$  is often referred to as "the hessian matrix"
- Both the estimator and the hessian matrix are often found by numerical methods.





#### Likelihood for Beverton-Holt

• Here we will model the observations as:

$$\log R_i = \underbrace{\log(a) + \log(\mathsf{ssb}_i) - \log(1 + \exp(\log(b))\mathsf{ssb}_i)}_{\mu_i} + \varepsilon_i ,$$

where  $\varepsilon_i \sim N(0, \sigma^2)$  independent.

• or in other words:

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$$\log R_i \sim N(\mu_i, \sigma^2)$$
, independent.

• The negative log likelihood becomes:

$$\ell(\log(R_i)|\log(a), \log(b), \sigma^2) = \frac{n}{2}\log(2\pi\sigma^2) + \frac{1}{2\sigma^2}\sum_{i=1}^n (\log(R)_i - \mu_i)^2$$



## **Beverton-Holt in AD Model Builder**

```
DATA_SECTION
  init_int nR
  init_int nC
  init_matrix obs(1,nR,1,nC)
  vector ssb(1,nR)
  !! ssb=column(obs,1);
  vector logR(1,nR)
  !! logR=column(obs,2);
PARAMETER_SECTION
  init_number loga;
  init_number logb;
  init_number logSigma;
  sdreport_number sigmaSq;
  vector pred(1,nR);
  objective_function_value nll;
PROCEDURE_SECTION
  sigmaSq=exp(2.0*logSigma);
  pred=loga+log(ssb)-log(1+exp(logb)*ssb);
  nll=0.5*(nR*log(2*M_PI*sigmaSq)+sum(square(logR-pred))/sigmaSq);
```

The logarithm of the determinant of the hessian = 12.7374									
index	name	value	std dev	1	2	3	4		
1	loga	1.8684e+00	1.3963e-01	1.0000					
2	logb	-1.2055e+01	3.3381e-01	0.9357	1.0000				
3	logSigma	-1.0963e+00	1.0426e-01	-0.0000	0.0000	1.0000			
4	sigmaSq	1.1162e-01	2.3275e-02	-0.0000	0.0000	1.0000	1.0000		

## **Things to remember**

- This is just based on a normal approximation, so it is not always appropriate. Also we may want to compute our C.I. on log scale and then transfer the quantiles the original scale to ensure that a parameter is positive.
- Variables other than the actual model parameters can be added to the output reports by declaring them as sdreport\_ in the parameter section (see e.g. sigmaSq)
- The three usual output files are called **\*.par**, **\*.std**, and **\*.cor**





## **Profile likelihood**

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• Profile likelihood for a given parameter is easily accessible in AD Model Builder

```
DATA_SECTION
  init_int nR
  init int nC
  init_matrix obs(1,nR,1,nC)
  vector ssb(1,nR)
  !! ssb=column(obs,1);
  vector logR(1,nR)
  !! logR=column(obs,2);
PARAMETER SECTION
  init_number loga;
  init_number logb;
  init_number logSigma;
  sdreport_number sigmaSq;
  vector pred(1,nR);
  likeprof_number loga_pl
                                 // NOTICE
  likeprof_number logb_pl
                                 // NOTICE
  objective_function_value nll;
PROCEDURE_SECTION
  loga_pl=loga;
                                 //NOTICE
  logb_pl=logb;
                                 //NOTICE
  sigmaSq=exp(2.0*logSigma);
  pred=loga+log(ssb)-log(1+exp(logb)*ssb);
  nll=0.5*(nR*log(2*M_PI*sigmaSq)+sum(square(logR-pred))/sigmaSq);
```

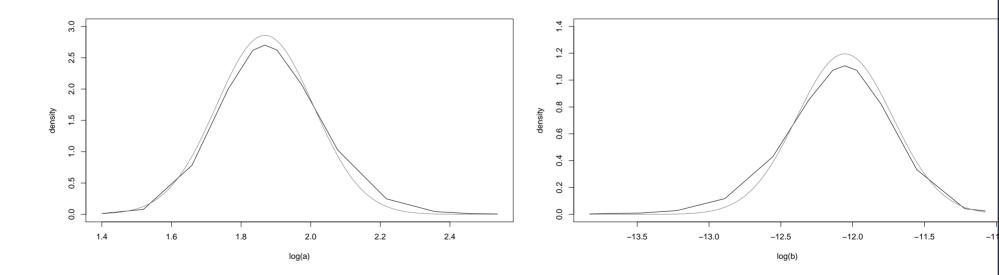
- Running the program with ./modelname -lprof
- Will produce two files loga\_pl.plt and loga\_pl.plt



## Profile likelihood (cont.)

- Each of those files have (x,y) points of the profile density
  - loga\_pl: Profile likelihood 1.39962 0.013531 1.41957 0.024403 1.43951 0.035275 1.45946 0.046147 1.47941 0.057019 1.49935 0.067891 1.5193 0.078763 1.53925 0.179453

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• Options are available to make more support points.

## What is MCMC and which variant are we using

- Assume we have an unnormalized probability density function  $\phi(\theta)$
- MCMC is a collection of methods to simulate a Markov chain  $\theta_1, ..., \theta_N$  with an equilibrium distribution given by  $\phi(\theta)$
- This is probably known to some from WinBUGS
- AD Model Builder uses what is known as a RW-MH (Random Walk Metropolis-Hastings)
- The starting point is  $\hat{\theta}$  and the proposal variance is  $\mathrm{var}(\hat{\theta})$





## **Example: The negative binomial**

• Assume that these 15 numbers follow a negative binomial distribution:

```
13 \ 5 \ 28 \ 28 \ 15 \ 4 \ 13 \ 4 \ 10 \ 17 \ 11 \ 13 \ 12 \ 17 \ 3
```

• The model is coded as:

```
DATA_SECTION
int N
!! N=15;
init_vector X(1,N)
PARAMETER_SECTION
init_number logsize;
init_bounded_number p(0,1);
sdreport_number size;
sdreport_number size;
sdreport_number pp;
objective_function_value nll;
PROCEDURE_SECTION
size=exp(logsize);
pp=p;
nll=-sum(gammln(X+size))+N*gammln(size)+
sum(gammln(X+1.0))-N*size*log(p)-sum(X)*log(1.0-p);
```

index	name	value	std dev
1	logsize	1.3017e+00	4.7101e-01
2	р	2.2218e-01	8.5571e-02
3	size	3.6754e+00	1.7312e+00
4	рр	2.2218e-01	8.5571e-02

#### **Basic use**

• Simply run the model with -mcmc N, where N is the number of steps. For instance:

an@ch-pcb-an:~\$./simplenbin -mcmc 10000

 The file <modelname>.hst then contains points on the simulated pdf of all sdreport variables.
 # samples sizes

```
10000
# step size scaling factor
1.2
# step sizes
 0.549441 0.0223384
# means
 4.33401 0.242679
# standard devs
 4.39827 0.178819
# lower bounds
 -8 -11
# upper bounds
 34 19
#number of parameters
#current parameter values for mcmc restart
 0.826337 0.148048
#random nmber seed
1262173905
```





#size -0.0615169 0 0.487924 0.00127402 1.03736 0.0251165 ... 22.4655 0.000182003 23.015 0

#### #pp

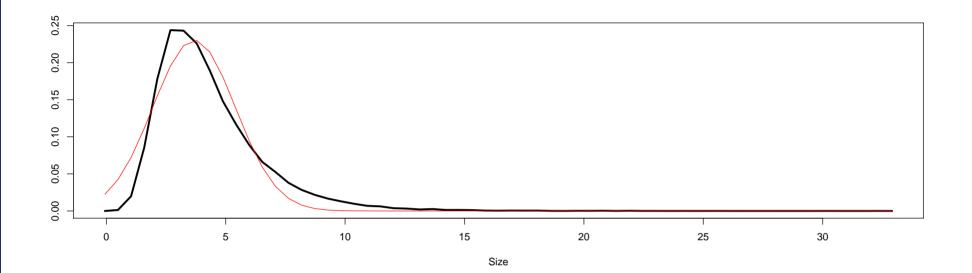
-0.00304368 0 0.0192947 0.00447659 0.0416332 0.0358127 ... 0.64477 0.00447659 0.667109 0

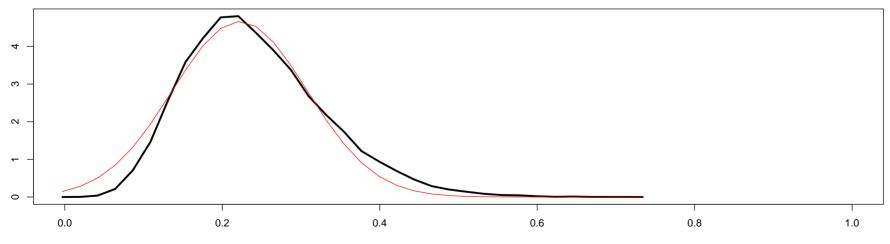




#### **Useful for plotting**

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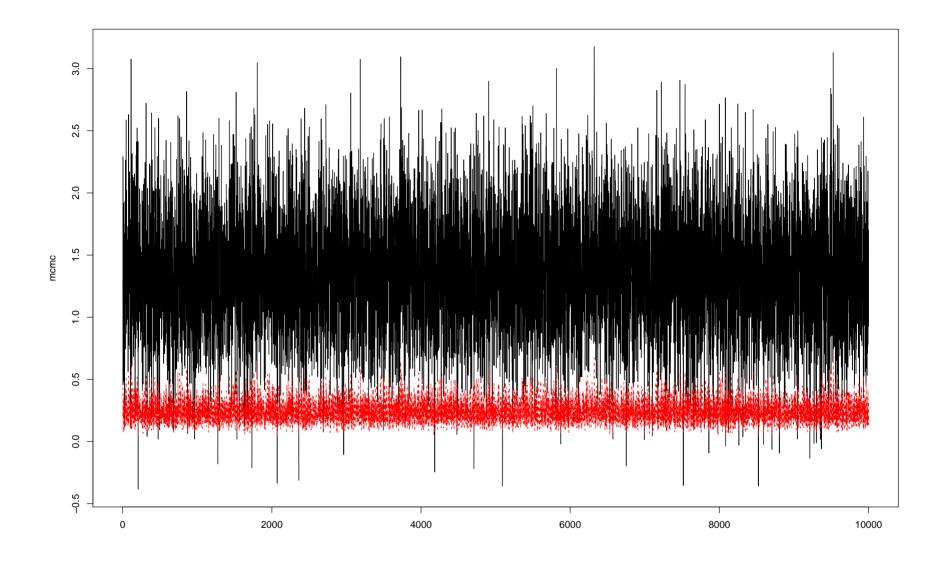
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#### Want to study the chain?

- The chain of parameters (not sdreport variables) can be saved by: an@ch-pcb-an:~\$./simplenbin -mcmc 100000 -mcsave 10
- here the -mcsave N tells it to save every N'th step
- Saves to a binary file <modelname>.psv, which can be read into R by:
  - > filen <- file("MCMC/simplenbin.psv", "rb")</pre>
  - > nopar <- readBin(filen, what = integer(), n = 1)</pre>
  - > mcmc <- readBin(filen, what = numeric(), n = nopar \* 10000)</pre>
  - > mcmc <- matrix(mcmc, byrow = TRUE, ncol = nopar)</pre>





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## The chain of custom output

- Suppose we want the output chain of something that is not a model parameter (here 'size')
- Then we need to change the code a bit

```
GLOBALS_SECTION
  #include <fstream.h>
  ofstream sizeout("size.cha");
DATA_SECTION
  int N
  !! N=15;
  init_vector X(1,N)
PARAMETER_SECTION
  init_number logsize;
  init_bounded_number p(0,1);
  sdreport_number size;
  sdreport_number pp;
  objective_function_value nll;
PROCEDURE_SECTION
  size=exp(logsize);
  pp=p;
  nll=-sum(gammln(X+size))+N*gammln(size)+
       sum(gammln(X+1.0))-N*size*log(p)-sum(X)*log(1.0-p);
  if(mceval_phase()){
    ofstream sizeout("size.cha", ios::app);
    sizeout<<size<<"\n";</pre>
  }
```



• To run we must type:

an@ch-pcb-an:~\$./simplenbin -mcmc 100000 -mcsave 10
an@ch-pcb-an:~\$./simplenbin -mceval

- $y = \left( \begin{array}{c} y = \left( \end{array}{y = \left( \begin{array}{c} y = \left( \end{array}{y = \left( \begin{array}{c} y = \left( \end{array}{y =$
- And then **Size**. cha is produced





# **Exercise: Beverton-Holt comparing sequential Bayesi**an and integrated analysis

- The results shown in the Beverton-Holt examples were based only on the data file bh.dat
- For a similar stock in a similar area we have these estimates
  - 1loga1.8085e+001.2725e-011.00002logb-1.2183e+013.2431e-010.92781.0000
- Modify the program **bh.tpl** to use this as prior information.
- In the file bh0.dat we have the entire data set from the similar stock.
- Modify the program **bh.tpl** to use the both data sets to estimate log(a) and log(b).
- Use MCMC and plot the joint distribution of log(a) and log(b). Make one plot for the sequential Bayesian approach, and one for the integrated.





## Solution

• To include the estimates and their estimated covariance as prior information the code is modified as:

```
DATA_SECTION
  init_int nR
  init int nC
  init_matrix obs(1,nR,1,nC)
  vector ssb(1,nR)
  !! ssb=column(obs,1);
  vector logR(1,nR)
  !! logR=column(obs,2);
  init_vector priorMean(1,2)
  init_vector priorSd(1,2)
  init_matrix priorCor(1,2,1,2)
  matrix priorCov(1,2,1,2)
  !! priorCov=elem_prod(outer_prod(priorSd, priorSd), priorCor);
PARAMETER_SECTION
  init_number loga;
  init_number logb;
  init_number logSigma;
  sdreport_number sigmaSq;
  vector pred(1,nR);
  vector logab(1,2);
  vector diff(1,2);
  objective_function_value nll;
PROCEDURE_SECTION
  sigmaSq=exp(2.0*logSigma);
  pred=loga+log(ssb)-log(1+exp(logb)*ssb);
  nll=0.5*(nR*log(2*M_PI*sigmaSq)+sum(square(logR-pred))/sigmaSq);
  logab(1)=loga; logab(2)=logb;
  diff=logab-priorMean;
```

```
nll+=0.5*(log(2.0*M_PI)*2.0+log(det(priorCov))+diff*inv(priorCov)*diff);
```



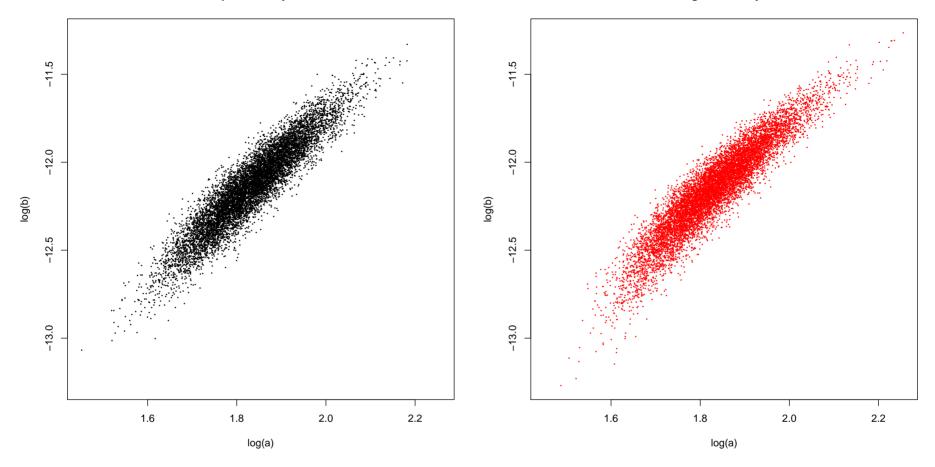


• To use both data sets the code is modified as:

```
DATA SECTION
  init_int nR
  init_int nC
  init_matrix obs(1,nR,1,nC)
  vector ssb(1,nR)
  !! ssb=column(obs,1);
  vector logR(1,nR)
  !! logR=column(obs,2);
  !! ad_comm::change_datafile_name("bh0.dat");
  init_int nR0
  init_int nCO
  init_matrix obs0(1,nR0,1,nC0)
  vector ssb0(1,nR0)
  !! ssb0=column(obs0,1);
  vector logR0(1,nR0)
  !! logR0=column(obs0,2);
PARAMETER_SECTION
  init_number loga;
  init_number logb;
  init_number logSigma;
  sdreport_number sigmaSq;
  vector pred(1,nR);
  init_number logSigma0;
  sdreport_number sigmaSq0;
  vector pred0(1,nR0);
  objective_function_value nll;
PROCEDURE_SECTION
  sigmaSq=exp(2.0*logSigma);
  pred=loga+log(ssb)-log(1+exp(logb)*ssb);
  nll=0.5*(nR*log(2*M_PI*sigmaSq)+sum(square(logR-pred))/sigmaSq);
  sigmaSq0=exp(2.0*logSigma0);
  pred0=loga+log(ssb0)-log(1+exp(logb)*ssb0);
  nll+=0.5*(nR0*log(2*M_PI*sigmaSq0)+sum(square(logR0-pred0))/sigmaSq0);
```

Sequential Bayesian

Integrated analysis



• Identical?

